

A CHARACTERIZATION OF THE POWER FUNCTION DISTRIBUTION BY INDEPENDENT PROPERTY OF LOWER RECORD VALUES

EUN-HYUK LIM* AND MIN-YOUNG LEE**

ABSTRACT. We prove a characterization of the power function distribution by lower record values. We prove that $F(x) = \left(\frac{x}{a}\right)^\alpha$ for all x , $0 < x < a$, $\alpha > 0$ and $a > 0$ if and only if $\frac{X_{L(n)}}{X_{L(m)}}$ and $X_{L(m)}$ are independent for $1 \leq m < n$.

1. Introduction

Let $\{X_n, n \geq 1\}$ be a sequence of independent and identically distributed (i.i.d.) random variables with cumulative distribution function (cdf) $F(x)$ and probability density function (pdf) $f(x)$. Let $Y_n = \max(\min)\{X_1, X_2, \dots, X_n\}$ for $n \geq 1$. We say X_j is an upper(lower) record value of this sequence, if $Y_j > (<)Y_{j-1}$ for $j > 1$. We denote by $X_{U(m)}$ and $X_{L(m)}$ the m -th upper and lower record values, respectively. By definition, X_1 is an upper as well as a lower record value. The indices at which the upper record values occur are given by the record times $\{U(n), n \geq 1\}$, where $U(n) = \min\{j \mid j > U(n-1), X_j > X_{U(n-1)}, n \geq 2\}$ with $U(1) = 1$. We assume that all upper record values $X_{U(i)}$ for $i \geq 1$ occur at a sequence $\{X_n, n \geq 1\}$ of i.i.d. random variables.

We define the power function distribution of a random variable.

A continuous random variable X is called the power function distribution with parameters $a > 0$, $\alpha > 0$ if its cdf is given by

Received September 24, 2012; Accepted January 11, 2013.

2010 Mathematics Subject Classification: Primary 62E10, 62E50.

Key words and phrases: power function distribution, independent and identically distributed, lower record values.

Correspondence should be addressed to Min-Young Lee, leemy@dankook.ac.kr.

$$(1.1) \quad F(x) = \begin{cases} \left(\frac{x}{a}\right)^\alpha, & 0 < x < a, \\ 0, & \text{otherwise.} \end{cases}$$

Ahsanullah(1995) prove that X has exponential distribution if and only if $X_{U(n)} - X_{U(m)}$ and $X_{U(m)}$ are independent. Also, Lee and Lim(2010) show that X has weibull distribution if and only if $\frac{X_{U(m)}}{X_{U(n)}}$ and $X_{U(n)}$ are independent.

In this paper, we obtain a characterization of the power function distribution by independent property of lower record values.

2. Main results

To prove Theorem 2.2, we need the following Lemma 2.1.

LEMMA 2.1. *Let $F(x)$ be an absolutely continuous function and $F(x) > 0$ for all $x > 0$. Suppose that $\frac{F(vw)}{F(v)} = e^{-q(v,w)}$ and $h(v, w) = \{q(v, w)\}^r e^{-q(v,w)} \{-\frac{\partial}{\partial w} q(v, w)\}$ for $r \geq 0$, $h(v, w) \neq 0$, $\frac{\partial}{\partial w} q(v, w) \neq 0$ for any v and w . If $h(v, w)$ is independent of v , then $q(v, w)$ is a function of w only.*

Proof. Let

$$(2.1) \quad \begin{aligned} g(w) &= h(v, w) = \{q(v, w)\}^r e^{-q(v,w)} \{-\frac{\partial}{\partial w} q(v, w)\} \\ &= \sum_{j=0}^{\infty} \frac{(-1)^j}{j!} q(v, w)^{j+r} \{-\frac{\partial}{\partial w} q(v, w)\}. \end{aligned}$$

Integrating (2.1) with respect to w , we obtain

$$(2.2) \quad \int g(w)dw + c = \sum_{j=0}^{\infty} \frac{(-1)^{j+1}}{j!} q(v, w)^{j+r+1} \frac{1}{(j+r+1)} = G_1(w).$$

Here G_1 is a function of w only and c is independent of w but may depend on v .

Now letting $w \rightarrow 1$, $q(v, w) \rightarrow 0$, we have c independently of v from (2.2). Therefore

$$\begin{aligned} 0 &= \frac{\partial}{\partial v} G_1(w) = \sum_{j=0}^{\infty} \frac{(-1)^j}{j!} q(v, w)^{j+r} \left\{ -\frac{\partial}{\partial v} q(v, w) \right\} \\ &= g(w) \left\{ -\frac{\partial}{\partial w} q(v, w) \right\}^{-1} \left\{ \frac{\partial}{\partial v} q(v, w) \right\}. \end{aligned}$$

We know $h(v, w) \neq 0$ and $\frac{\partial}{\partial w} q(v, w) \neq 0$, so we must have

$$\frac{\partial}{\partial v} q(v, w) = 0.$$

Hence $q(v, w)$ is a function of w only. □

THEOREM 2.2. *Let $\{X_n, n \geq 1\}$ be a sequence of i.i.d. random variables with cdf $F(x)$ which is absolutely continuous with pdf $f(x)$ and $F(a) = 1$ and $F(x) < 1$ for all $x, 0 < x < a$. Then $F(x) = (\frac{x}{a})^\alpha$ for all $x, 0 < x < a$ and $a > 0$ if and only if $\frac{X_{L(n)}}{X_{L(m)}}$ and $X_{L(m)}$ are independent for $1 \leq m < n$.*

Proof. The joint pdf $f_{m,n}(x, y)$ of $X_{L(m)}$ and $X_{L(n)}$ is found to be

$$f_{m,n}(x, y) = \frac{\{H(x)\}^{m-1}}{(m-1)!} h(x) \frac{\{H(y) - H(x)\}^{n-m-1}}{(n-m-1)!} f(y),$$

where $H(x) = -\ln F(x)$ and $h(x) = -\frac{d}{dx} H(x)$.

Consider the functions $V = X_{L(m)}$ and $W = \frac{X_{L(n)}}{X_{L(m)}}$. It follows that $x_{L(m)} = v, x_{L(n)} = vw$ and $|J| = v$. Thus we can find the joint pdf $f_{V,W}(v, w)$ of V and W as

$$f_{V,W}(v, w) = \frac{\{H(v)\}^{m-1}}{(m-1)!} h(v) \frac{\{H(vw) - H(v)\}^{n-m-1}}{(n-m-1)!} f(vw)v$$

for $0 < v < a, 0 < w < 1$.

If $F(x) = (\frac{x}{a})^\alpha$ for all $0 < x < a$ and $a > 0$, then we get

$$\begin{aligned} (2.3) \quad f_{V,W}(v, w) &= \frac{\alpha^2}{a(m-1)!(n-m-1)!} \\ &\quad \times \left\{ -\alpha \ln \left(\frac{v}{a}\right) \right\}^{m-1} \left\{ -\alpha \ln w \right\}^{n-m-1} \left(\frac{v}{a}\right)^{\alpha-1} w^{\alpha-1} \end{aligned}$$

for all $v > 1, w > 1$ and $\alpha > 0$.

The marginal pdf of W is given by

$$(2.4) \quad f_W(w) = \int_0^a f_{V,W}(v, w) dv = \alpha \frac{\{-\alpha \ln w\}^{n-m-1}}{(n-m-1)!} w^{\alpha-1}$$

for all $0 < w < 1$, $\alpha > 0$.

Also, the pdf $f_V(v)$ is given by

$$(2.5) \quad \begin{aligned} f_V(v) &= \frac{\{H(v)\}^{n-2}}{(n-2)!} f(v) \\ &= \frac{\alpha}{a(m-1)!} \left\{-\alpha \ln \left(\frac{v}{a}\right)\right\}^{m-1} \left(\frac{v}{a}\right)^{\alpha-1}. \end{aligned}$$

From (2.3), (2.4), and (2.5), we obtain $f_{V,W}(v, w) = f_V(v)f_W(w)$. Hence V and W are independent for $1 < m < n$.

Now we will prove the sufficient condition. Let us use the transformation $V = X_{L(m)}$ and $W = \frac{X_{L(n)}}{X_{L(m)}}$. The Jacobian of the transformation is $|J| = v$. Thus we can find the joint pdf $f_{V,W}(v, w)$ of V and W as

$$(2.6) \quad f_{V,W}(v, w) = \frac{\{H(v)\}^{m-1}}{(m-1)!} h(v) \frac{\{H(vw) - H(v)\}^{n-m-1}}{(n-m-1)!} f(vw)v$$

for all $0 < v < a$, $0 < w < 1$ and $\alpha > 0$.

The pdf $f_V(v)$ is given by

$$(2.7) \quad f_V(v) = \frac{\{H(v)\}^{m-1}}{(m-1)!} f(v)$$

for all $0 < v < a$, $m > 1$.

From (2.6) and (2.7), we can get the conditional pdf of $f_W(w|v)$ as

$$\begin{aligned} f_W(w|X_{L(m)} = v) &= \frac{\{H(vw) - H(v)\}^{n-m-1} f(vw)v}{(n-m-1)! F(v)} \\ &= \frac{1}{(n-m-1)!} \left(-\ln \frac{F(vw)}{F(v)}\right)^{n-m-1} \frac{f(vw)v}{F(v)} \\ &= \frac{1}{(n-m-1)!} \left(-\ln \frac{F(vw)}{F(v)}\right)^{n-m-1} \left(\frac{F(vw)}{F(v)}\right) \left(-\frac{\partial}{\partial w} \left(-\ln \frac{F(vw)}{F(v)}\right)\right). \end{aligned}$$

Since V and W are independent, by using Lemma 2.1, $q(v, w) = -\ln \frac{F(vw)}{F(v)}$ is a function of w only. Thus

$$\frac{F(vw)}{F(v)} = G(w),$$

where $G(w)$ is a function of w only. Taking $v \rightarrow a-$, we get $G(w) = F(aw)$. Thus

$$(2.8) \quad F(vw) = F(v)F(aw)$$

for all $0 < v < a$ and $0 < w < 1$.

By the theory of functional equations, the only continuous solution of (2.8) with the boundary condition $F(0) = 0$ is

$$F(x) = \left(\frac{x}{a}\right)^\alpha$$

for all $0 < x < a$ and $\alpha > 0$.

This completes the proof. \square

References

- [1] J. Aczel, *Lectures on Functional Equation and Their Applications*, Academic Press, Newyork, 1966.
- [2] M. Ahsanullah, *Record Statistics*, Nova science Publishers, Inc. NJ, USA, 1995.
- [3] M. Ahsanullah and Z. Raqab, *Bounds and Characterizations of Record Statistics*, Nova science Publishers, Inc, 2006.
- [4] M. Y. Lee and E. H. Lim, *On Characterizations of the Weibull distribution by the independent property of record values*, J. Chungcheong Math. Soc. **23** (2010), no. 2, 245-250.

*

Department of Mathematics
Dankook University
Cheonan 330-714, Republic of Korea
E-mail: ehlim@dankook.ac.kr

**

Department of Mathematics
Dankook University
Cheonan 330-714, Republic of Korea
E-mail: leemy@dankook.ac.kr